

# Stochastic Optimization Without Integration

João P. Hespanha

Center for Control  
Dynamical Systems and Computation



UC SANTA BARBARA  
**engineering**



Problem addressed:

$$\min_u E_d [f(u, d)]$$

- no closed form for expected value
- willing to sacrifice accuracy for speed
- error bounds

Talk outline:

1. Motivating problems
2. Additive/multiplicative bounds for expected values
3. Bounds based on Laplace method for integration

*In collaboration with:*

*Raphael Chinchilla, Murat Erdal, Dr. Guosong Yang (UCSB)*

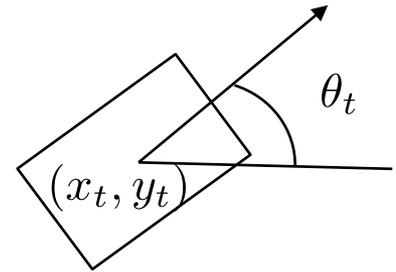
*Prof. Ramon Costa (Federal Univ. Rio Janeiro, Brazil)*

*Prof. Kevin Plaxco (UCSB)*

Dubins vehicle (discrete-time)

$$\begin{aligned}x_{t+1} &= x_t + v \left( \cos \theta_t - \frac{u_t}{2} \sin \theta_t \right) + d_t \\y_{t+1} &= y_t + v \left( \sin \theta_t + \frac{u_t}{2} \cos \theta_t \right) + w_t \\ \theta_{t+1} &= \theta_t + u_t + \nu_t\end{aligned}$$

angular velocity (control)  
Cartesian disturbances  
rotational disturbance



$$\min_{u_1, \dots, u_T} \mathbb{E} \left[ \sum_{t=1}^T x_t^2 + y_t^2 + u_t^2 \right]$$

- not possible to compute expectation in closed form, even if  $d_t, w_t, \nu_t$  multivariable Gaussian/von Misses

yet...

- need to solve fast for receding-horizon / Model Predictive Control (MPC)

Dynamical system:

$$\begin{aligned}x_{t+1} &= f(x_t; \theta) + d_t \leftarrow \text{state disturbances} \\y_t &= g(x_t; \theta) + w_t \leftarrow \text{measurement noise}\end{aligned}$$

$\theta$   $\leftarrow$  unknown parameter

Maximum likelihood estimation:

$$\hat{\theta}_{\text{MLE}} = \arg \max_{\theta} p(y_1, \dots, y_K; \theta)$$

$p(y_1, \dots, y_K; \theta)$   $\leftarrow$  prob. density function (pdf) of output

Dynamical system:

$$\begin{aligned}x_{t+1} &= f(x_t; \theta) + d_t \\y_t &= g(x_t; \theta) + w_t\end{aligned}$$

state disturbances  
measurement noise  
unknown parameter

Maximum likelihood estimation:

$$\hat{\theta}_{\text{MLE}} = \arg \max_{\theta} p(y_1, \dots, y_K; \theta)$$

prob. density function (pdf) of output

- output pdf is generally difficult to compute, but
- **conditional** output pdf **given state** is typically very easy to compute
- can write MLE as

$$\hat{\theta}_{\text{MLE}} = \arg \max_{\theta} \mathbb{E} \left[ p(y_1, \dots, y_K \mid x_1, \dots, x_K; \theta) \right]$$

law of total expectation  
expectation w.r.t.  $x_t$

- generally no closed-form expression for expectation yet,...
- need to solve fast to compute  $\theta$ -dependent a-posterior estimate of state, given measurements for output-feedback control

Dynamical system:

$$x_{t+1} = A(\theta)x_t + B(\theta)u_t + d_t \quad \leftarrow \text{state disturbances}$$

$$y_t = C(\theta)x_t + w_t \quad \leftarrow \text{measurement noise}$$

$\nwarrow$  unknown parameter

Goal: select  $u_1, \dots, u_T$  to minimize error for estimator  $\hat{\theta}(y_1, y_2, \dots, y_T)$

# Experiment Design

Dynamical system:

$$\begin{aligned}x_{t+1} &= A(\theta)x_t + B(\theta)u_t + d_t \quad \leftarrow \text{state disturbances} \\y_t &= C(\theta)x_t + w_t \quad \leftarrow \text{measurement noise}\end{aligned}$$

$\theta$  unknown parameter

Goal: select  $u_1, \dots, u_T$  to minimize error for estimator  $\hat{\theta}(y_1, y_2, \dots, y_T)$

Assuming unbiased estimator that achieves Cramer-Rao lower bound

$$\mathbb{E} \left[ (\hat{\theta} - \theta)(\hat{\theta} - \theta)' \right] = \text{FIM}(\theta)^{-1}$$

$\text{FIM}(\theta) := -\mathbb{E} \left[ \frac{\partial^2 \log p(y_1, \dots, y_T; \theta)}{\partial \theta^2} \right]$

Fisher information matrix      w.r.t noise/disturbances      Hessian matrix of measurements pdf

*A-optimality:*

$$\min_{u_1, \dots, u_T} \mathbb{E} \left[ \text{trace FIM}(\theta)^{-1} \right]$$

w.r.t  $\theta$       minimizes parameter estimates MSE

*D-optimality:*

$$\min_{u_1, \dots, u_T} \mathbb{E} \left[ \log \det \text{FIM}(\theta)^{-1} \right]$$

w.r.t  $\theta$       minimizes volume of parameters estimates confidence ellipsoids

Monte Carlo Integration:

$$\mathbb{E}[f(d)] \approx \frac{1}{K} \sum_{k=1}^K f(d_k) \quad \text{error} \propto O\left(\frac{\sigma}{K^{\frac{1}{2}}}\right)$$

random variable

indep. identically distributed (iid) samples of  $d$

- central limit theorem
- dimension-independent

In stochastic optimization, leads to Sample Average Approximation (SAA):

$$\min_u \mathbb{E}[f(u, d)] \approx \min_u \frac{1}{K} \sum_{k=1}^K f(u, d_k)$$

random variable

error  $\propto \begin{cases} O\left(\frac{1}{K}\right) & \text{strongly convex and Lipschitz} \\ O\left(\frac{1}{K^{\frac{1}{2}}}\right) & \text{Lipschitz} \end{cases}$

[Shalev-Shwartz et al. 2010]

[Kleywegt et al, 2001]

e.g., stochastic gradient descent:

$$u(i+1) = u(i) - \gamma(i) \frac{1}{K} \sum_{k=1}^K \nabla f_u(u(i), d_k)$$

step size

gradient of empirical average

## MC-based methods

- can be made arbitrarily accurate by making  $K \rightarrow \infty$
- generally slow (need large  $K$ , # iterations)
- often computationally expensive to sample from desired distribution (especially for conditional distribution given measurements, typically requiring Markov Chain Monte Carlos method, e.g., Metropolis Hastings or Gibbs sampling)

## Our goal...

- Trade accuracy for speed
- If possible, get bounds on error

[Shalev-Shwartz et al. 2010]

random  
variable

$$\text{error} \propto \begin{cases} O\left(\frac{1}{K}\right) & \text{strongly convex and Lipschitz} \\ O\left(\frac{1}{K^{\frac{1}{2}}}\right) & \text{Lipschitz} \end{cases}$$

[Kleywegt et al, 2001]

e.g., stochastic gradient descent:

$$u(i+1) = u(i) - \gamma(i) \frac{1}{K} \sum_{k=1}^K \nabla f_u(u(i), d_k)$$

step size

gradient of empirical average

Problem addressed:

$$\min_u E_d [f(u, d)]$$

- no closed form for expected value
- willing to sacrifice accuracy for speed
- error bounds

Talk outline:

1. Motivating problems
2. Additive/multiplicative bounds for expected values
3. Bounds based on Laplace method for integration

*In collaboration with:*

*Raphael Chinchilla, Murat Erdal, Dr. Guosong Yang (UCSB)*

*Prof. Ramon Costa (Federal Univ. Rio Janeiro, Brazil)*

*Prof. Kevin Plaxco (UCSB)*

# Coarse Bounds on Expectations

Suppose  $d$  is random variable taking values on  $\mathcal{D}$ , and

$$f_{\min} \leq f(d) \leq f_{\max} \quad \text{with probability 1}$$

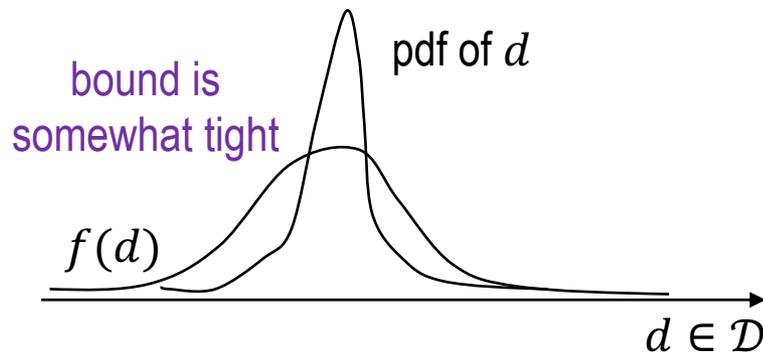
by monotonicity

$$f_{\min} \leq \mathbb{E}[f(d)] \leq f_{\max}$$

tightest such (upper) bound

$$\mathbb{E}[f(d)] \leq \underbrace{\text{ess sup}_{d \in \mathcal{D}} f(d)}_{\text{can discount zero-measure sets}}$$

values of  $d$  with positive but small probability increase supremum but have little impact on expectation



# Coarse Bounds on Expectations

Suppose  $d$  is random variable taking values on  $\mathcal{D}$ , and

$$f_{\min} \leq f(d) \leq f_{\max} \quad \text{with probability 1}$$

by monotonicity

$$f_{\min} \leq \mathbb{E}[f(d)] \leq f_{\max}$$

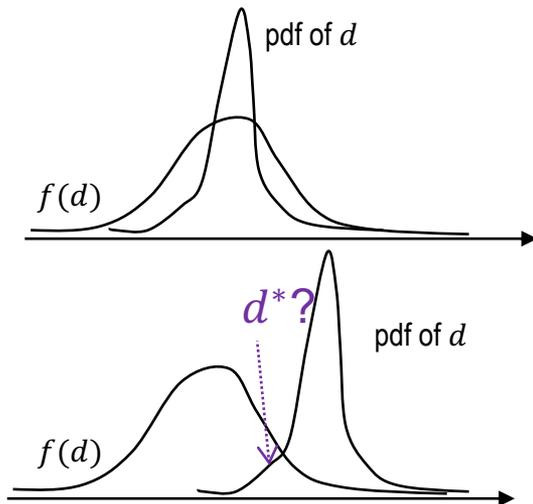
tightest such (upper) bound

$$\mathbb{E}[f(d)] \leq \text{ess sup}_{d \in \mathcal{D}} f(d)$$

can discount zero-measure sets

“naïf” idea:

$$\mathbb{E}[f(d)] \leq f(d^*)$$



prob. density function (pdf) of  $d$

$$d^* = \arg \max_{d \in \mathcal{D}} \left( f(d) + \log p(d) \right)$$

max will not pick value for  $d$  with low probability  
(very negative log-pdf)

# A Precise Bound

Additive bound: For every  $\epsilon \geq 0$  (and assuming all expectations are finite),

$$\text{ess inf}_{d \in \mathcal{D}} \left( f(d) - \epsilon \log p(d) \right) - \epsilon H_d \leq \mathbb{E}[f(d)] \leq \text{ess sup}_{d \in \mathcal{D}} \left( f(d) + \epsilon \log p(d) \right) + \epsilon H_d$$

prob. density function (pdf) of  $d$

inf at value for which  $f(d)$  is small but pdf not very small

sup at value for which  $f(d)$  is large but pdf not very small

differential entropy of  $d$   
 $H_d := \mathbb{E} \left[ -\log p(d) \right]$

# A Precise Bound

Additive bound: For every  $\epsilon \geq 0$  (and assuming all expectations are finite),

$$\text{ess inf}_{d \in \mathcal{D}} \left( f(d) - \epsilon \log p(d) \right) - \epsilon H_d \leq \mathbb{E}[f(d)] \leq \text{ess sup}_{d \in \mathcal{D}} \left( f(d) + \epsilon \log p(d) \right) + \epsilon H_d$$

prob. density function (pdf) of  $d$

inf at value for which  $f(d)$  is small but pdf not very small

sup at value for which  $f(d)$  is large but pdf not very small

differential entropy of  $d$   
 $H_d := \mathbb{E} \left[ -\log p(d) \right]$

Why? (upper bound)

$$\begin{aligned} \mathbb{E}[f(d)] &= \mathbb{E} \left[ f(d) + \epsilon \log p(d) - \epsilon \log p(d) \right] \\ &= \mathbb{E} \left[ f(d) + \epsilon \log p(d) \right] + \epsilon H_d \end{aligned}$$

use "trivial" bound on this term

$$\mathbb{E} \left[ f(d) + \epsilon \log p(d) \right] \leq \text{ess sup}_{d \in \mathcal{D}} f(d) + \epsilon \log p(d)$$

bound typically tighter as probability mass concentrated around large values of  $f(d) + \epsilon \log p(d)$

# More Bounds

Additive bound: For every  $\epsilon \geq 0$  (and assuming all expectations are finite),

$$\text{ess inf}_{d \in \mathcal{D}} \left( f(d) - \epsilon \log p(d) \right) - \epsilon H_d \leq \mathbb{E}[f(d)] \leq \text{ess sup}_{d \in \mathcal{D}} \left( f(d) + \epsilon \log p(d) \right) + \epsilon H_d$$

prob. density function (pdf) of  $d$   
 $\swarrow$   
 differential entropy of  $d$   
 $H_d := \mathbb{E} \left[ -\log p(d) \right]$

Multiplicative bound: For every  $\epsilon > 0$  (and assuming all expectations are finite),

$$\text{ess inf}_{d \in \mathcal{D}} \left( f(d) p(d)^{-\epsilon} \right) I_d(-\epsilon) \leq \mathbb{E}[f(d)] \leq \text{ess sup}_{d \in \mathcal{D}} \left( f(d) p(d)^\epsilon \right) I_d(\epsilon)$$

$I_d(\epsilon) := \mathbb{E} \left[ p(d)^{-\epsilon} \right]$

More generally: For every function  $\alpha: \mathcal{D} \rightarrow \mathbb{R}$

Any group operation that is

1. right ordered  $a \leq b \Leftrightarrow a \oplus c \leq b \oplus c$
2. E-distributive  $a \oplus \mathbb{E}[z] = \mathbb{E}[a \oplus z]$

$$\mathbb{E}[f(d)] \leq \text{ess sup}_{d \in \mathcal{D}} \left( f(d) \oplus \alpha(d) \right) \oplus A_d$$

for tight bound: pick  $\alpha, \oplus$  so that prob. mass concentrated around max. of  $f(d) \oplus \alpha(d)$

$$A_d := \mathbb{E} \left[ \neg \alpha(d) \right]$$

$$J^* := \min_u \left\{ \mathbb{E} [f(d, u)] : \mathbb{E} [g(d, u)] \leq 0 \right\}$$

replacing by upper bound will  
find “conservative/pessimistic”  
solution

replacing by upper bound will  
guarantee feasibility

or...

replacing both by lower bounds  
will find “optimistic” solution

E.g., using additive upper bounds:

$$J^\perp := \min_u \left\{ \text{ess sup}_{d \in \mathcal{D}} \left( f(d, u) + \epsilon \log p(d) \right) : g(\bar{d}, u) + \epsilon \log p(\bar{d}) \leq 0, \forall \bar{d} \in \mathcal{D} \right\}$$

1. arg-min guaranteed to be feasible for original problem & perform no worse than  $J^\perp$
2. cost  $J^\perp$  is an upper bound for  $J^*$
3. **replaced expectation/integration by optimization  $\Leftrightarrow$  min-max problem**
4. lower bound for  $J^*$  computable using lower bounds for expected values

# Application to Output-Feedback MPC

$$x_{t+1} = f(x_t, u_t; \theta) + d_t$$
$$y_t = g(x_t; \theta) + w_t$$

state disturbances

unknown parameter

measurement noise

$$\min_{u_0, \dots, u_{T-1}} \mathbb{E} \left[ \sum_{t=0}^T \ell(x_t, u_t) \mid y_{-K}, \dots, y_0 \right]$$

future controls

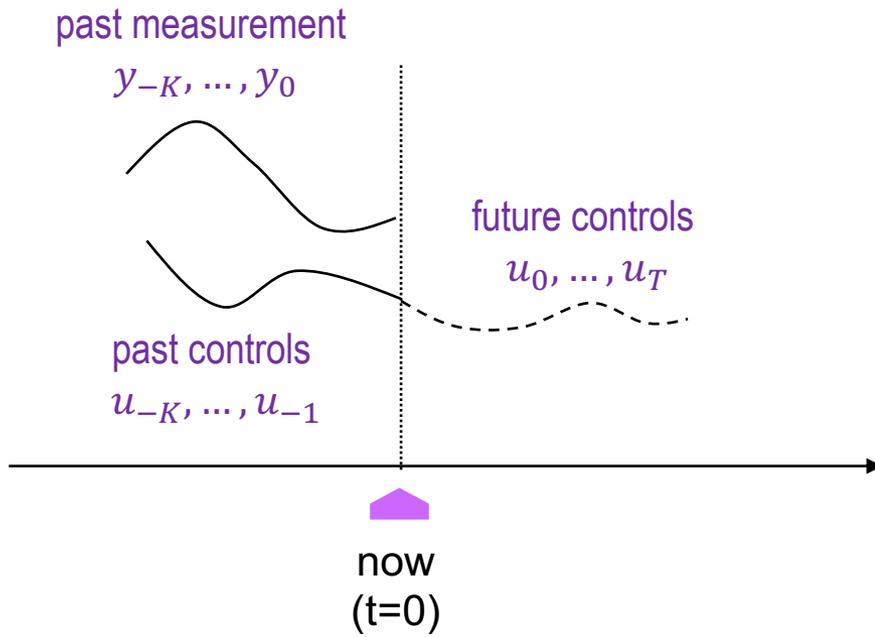
running cost, over future horizon

past measurements

a-posteriori (conditional) expectation with respect to:

1. unknown parameter  $\theta$
2. past noise  $w_{-K}, \dots, w_0$
3. past and future disturbances  $d_{-K}, \dots, d_{T-1}$

given past measurement  $y_{-K}, \dots, y_0$



In Model Predictive Control (MPC), optimization repeated at each time step under a receding horizon

$$x_{t+1} = f(x_t, u_t; \theta) + d_t$$

$$y_t = g(x_t; \theta) + w_t$$

$$\min_{u_0, \dots, u_{T-1}} \mathbb{E} \left[ \sum_{t=0}^T \ell(x_t, u_t) \mid y_{-K}, \dots, y_0 \right]$$

Observation: assuming  $x_{-K}, d_t, w_t, \theta$  independent, joint pdf of  $x_{-K}, d_t, y_t, \theta$  is easy to write

$$\left( \prod_{t=-K}^0 p_w(y_t - g(x_t; \theta)) \right) \left( \prod_{t=-K}^{T-1} p_d(x_{t+1} - f(x_t, u_t; \theta)) \right) p(x_{-K}) p(\theta)$$

pdfs of  $w_t, d_t, x_{-K}, \theta$

Why? Proof by induction to compute

$$p(x_{t+1}, y_t \mid d_{-K}, \dots, d_t, y_{-K}, \dots, y_{t-1}, x_{-K}, \theta)$$

starting at  $t = -K$  & using Bayes rule for induction step

Therefore,...

- closed-form solution for a-posteriori distribution (modulo normalization by pdf of  $y_{-K}, \dots, y_0$ )
- but not easy to draw samples from a-posteriori distribution (need Metropolis-Hastings or Gibbs sampling)

# Output-Feedback MPC

$$x_{t+1} = f(x_t, u_t; \theta) + d_t$$

$$y_t = g(x_t; \theta) + w_t$$

$$\min_{u_0, \dots, u_{T-1}} \mathbb{E} \left[ \sum_{t=0}^T \ell(x_t, u_t) \mid y_{-K}, \dots, y_0 \right]$$

Can be restated as:

$$J^* = \min_U \mathbb{E} [V(U, D) \mid Y]$$

$$U := \{u_0, \dots, u_{T-1}\}$$

$$D := \{x_{-K}, d_{-K}, \dots, d_{T-1}, \theta\}$$

$$Y := \{y_{-K}, \dots, y_0\}$$

Using additive bound:

$$J^\perp = \min_U \left( \text{ess sup}_D V(U, D) + \epsilon \log p(D|Y) \right) + \epsilon H_D$$

$$\log p(D|Y) = \log(D, Y) - \log P(Y)$$

$$H_D := \mathbb{E} [ -\log p(D)|Y ]$$

conditional pdf  
of D, given Y

joint pdf of  
D, given Y

marginal  
pdf of Y

diff. entropy of D

closed-form  
solution

indep. of U

indep. of U



$$\begin{aligned}x_{t+1} &= f(x_t, u_t; \theta) + d_t \\ y_t &= g(x_t; \theta) + w_t\end{aligned}$$

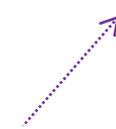
$$\min_{u_0, \dots, u_{T-1}} \mathbb{E} \left[ \sum_{t=0}^T \ell(x_t, u_t) \mid y_{-K}, \dots, y_0 \right]$$

Can be restated as:  $J^* = \min_U \mathbb{E} [V(U, D) \mid Y]$

$$\begin{aligned}U &:= \{u_0, \dots, u_{T-1}\} \\ D &:= \{d_{-K}, \dots, d_{T-1}, x_{-K}, \theta\} \\ Y &:= \{y_{-K}, \dots, y_0\}\end{aligned}$$

Using additive bound:

$$U^\perp = \arg \min_U \left( \text{ess sup}_D \sum_{t=1}^T \ell(x_t, u_t) + \epsilon \log p(D, Y) \right)$$

$$\left( \sum_{t=-K}^0 \log p_w(y_t - g(x_t; \theta)) \right) + \left( \sum_{t=-K}^{T-1} \log p_d(x_{t+1} - f(x_t, u_t; \theta)) \right) + \log p(x_{-K}) + \log p(\theta)$$


- all optimization terms computed in closed form without need for integration
- stochastic optimization replaced by min-max problem
- one-to-one map from  $d_t$ 's to  $x_t$ 's, so we can regard  $x_t$ 's as the optimization variables (computationally much better for 2<sup>nd</sup> order/Newton methods because of Hessian sparsity)
- this and all optimizations in remainder of paper solved using TensCalc toolbox

# LQG example

$$x_{t+1} = A(\theta)x_t + B(\theta)u_t + d_t$$

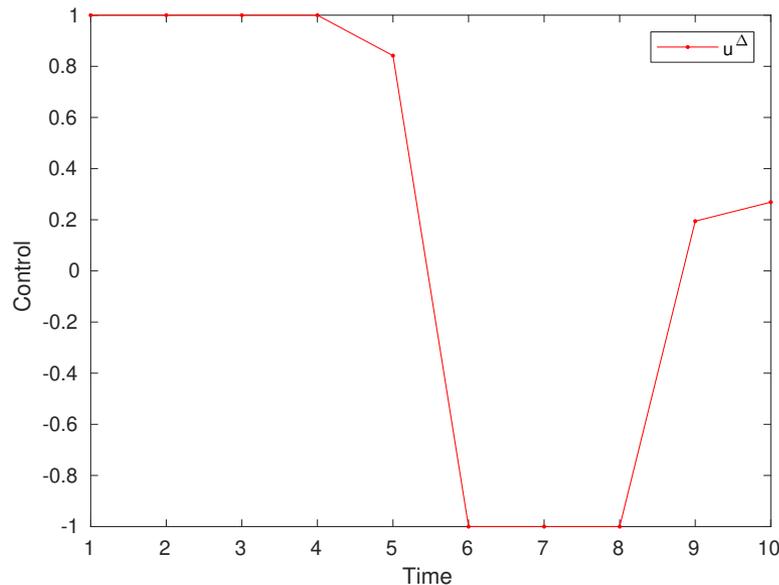
$$y_t = x_t + w_t$$

$$A(\theta) = \begin{bmatrix} \theta_1 & \theta_2 & 0 \\ 0 & \theta_3 & \theta_4 \\ 0 & 0 & \theta_5 \end{bmatrix} \quad B(\theta) = \begin{bmatrix} 0 \\ 0 \\ \theta_6 \end{bmatrix}$$

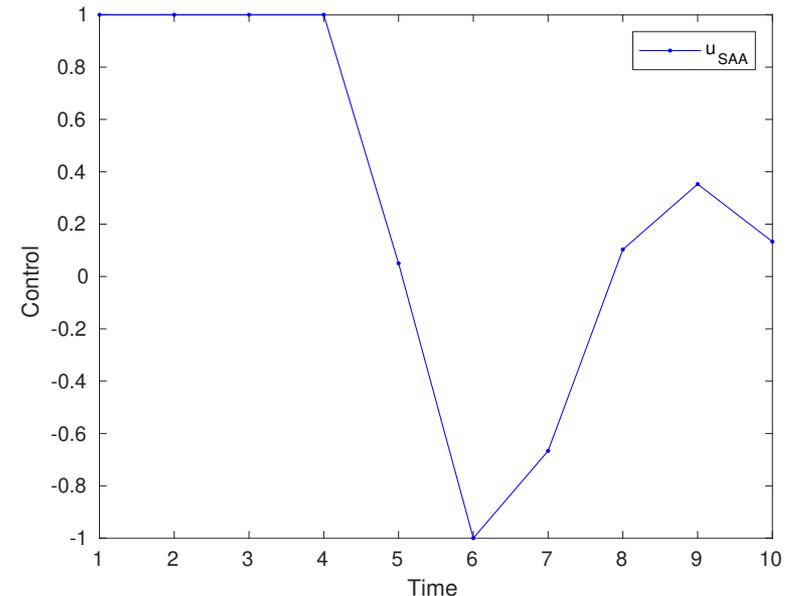
independent Gaussian parameters with

$$E[\theta_i] = 1, \text{Var}[\theta_i] = .5^2$$

$$T = 10, K = 20$$

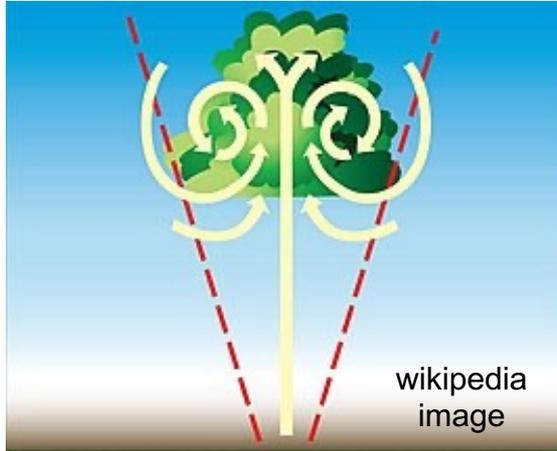


optimal control based on upper bound  
(criterion value 1.79e3, computation sub-second)



optimal control based SSA, with Gibbs sampling  
(criterion value 1.62e3, computation minutes)

thermal column of raising air



vertical wind speed at position  $x$

$$v(x) = v_0 e^{-\gamma \|x - p_0\|^2} + w$$

thermal parameters

stochastic variability

*Q: What is the best trajectory  $x_t$  to estimate thermal parameters  $\theta = [v_0, \gamma, p_0]$ , from point measurements of vertical wind speed?*





Problem addressed:

$$\min_u E_d [f(u, d)]$$

- no closed form for expected value
- willing to sacrifice accuracy for speed
- error bounds

Talk outline:

1. Motivating problems
2. Additive/multiplicative bounds for expected values
3. Bounds based on Laplace method for integration

*In collaboration with:*

*Raphael Chinchilla, Murat Erdal, Dr. Guosong Yang (UCSB)*

*Prof. Ramon Costa (Federal Univ. Rio Janeiro, Brazil)*

*Prof. Kevin Plaxco (UCSB)*

# Laplace Integration Method

$$\int_a^b e^{f(x)} dx \quad \text{smooth } f(x) \text{ with a maximum at } x_0 \text{ and } f''(x_0) < 0$$

Taylor series of  $f(x)$  around  $x_0$ :

$$f(x) = f(x_0) + \underbrace{f'(x_0)(x - x_0)}_{0 \text{ at maximum}} + \underbrace{\frac{f''(x_0)}{2}(x - x_0)^2}_{< 0} + O((x - x_0)^3)$$

Assuming 3<sup>rd</sup> and higher order terms negligible:

$$\int_a^b e^{f(x)} \approx \int_a^b e^{f(x_0) + \frac{f''(x_0)}{2}(x-x_0)^2} dx = e^{f(x_0)} \underbrace{\int_a^b e^{\frac{f''(x_0)}{2}(x-x_0)^2} dx}_{\text{Gaussian integral}}$$

Gaussian integral:

- closed-form for  $b = -a = \infty$

$$\int_{-\infty}^{\infty} e^{\frac{f''(x_0)}{2}(x-x_0)^2} dx = \sqrt{\frac{2\pi}{f''(x_0)}}$$

- in practice, integral is computed by finding maximum  $x_0$

marginal  
distribution

joint  
distribution

$$p(Y) = \int_{\mathbb{R}^n} p(Y, Z) dZ = \int_{\mathbb{R}^n} e^{\log p(Y, Z)} dZ$$

Taylor series of  $Z \mapsto \log p(Z, Y)$  around maximum  $Z_0$ :

$$\log p(Y, Z) = \log p(Y, Z_0) - \frac{1}{2} (Z - Z_0)' H(Y, Z_0) (Z - Z_0)' + O(\|Z - Z_0\|^3)$$
$$H(Y, Z_0) = - \left. \frac{\partial^2 \log p(Y, Z)}{\partial Z^2} \right|_{Z=Z_0}$$


# Laplace Integration Method

marginal  
distribution

joint  
distribution

$$p(Y) = \int_{\mathbb{R}^n} p(Y, Z) dZ = \int_{\mathbb{R}^n} e^{\log p(Y, Z)} dZ$$

Taylor series of  $Z \mapsto \log p(Z, Y)$  around maximum  $Z_0$ :

$$\log p(Y, Z) = \log p(Y, Z_0) - \frac{1}{2} (Z - Z_0)' H(Y, Z_0) (Z - Z_0) + O(\|Z - Z_0\|^3)$$

$H(Y, Z_0) = -\frac{\partial^2 \log p(Y, Z)}{\partial Z^2} \Big|_{Z=Z_0}$

Assuming 3<sup>rd</sup> and higher order terms negligible:

$$p(Y) \approx p(Y, Z_0) \underbrace{\int_{\mathbb{R}^n} e^{-\frac{1}{2} (Z - Z_0)' H(Y, Z_0) (Z - Z_0)} dZ}_{\text{Gaussian integral}} \propto \frac{p(Y, Z_0)}{\det H(Y, Z_0)}$$

joint distribution  
computed at maximum  $Z_0$

conditional  
means

$$E[Z|Y] = \int_{\mathbb{R}^n} Z e^{p(Z|Y)} dZ = \int_{\mathbb{R}^n} Z e^{p(Y, Z) - p(Y)} dZ \approx \dots = Z_0$$

$$\text{CoV}[Z|Y] = E[(Z - E[Z])(Z - E[Z])' | Y] \approx \dots = H(Y, Z_0)^{-1}$$

Given joint distribution  $p(Y, Z)$  we can compute marginal, conditional mean, and covariance matrix through optimization: typically easy to compute

$$Z_0 := \arg \max_Z \log p(Y, Z)$$

$$p(Y) \approx \frac{c p(Y, Z_0)}{\det H(Y, Z_0)}$$

useful for MLE

$$E[Z|Y] \approx Z_0$$

useful for Bayesian estimation

$$\text{CoV}[Z|Y] \approx \left. \frac{\partial^2 \log p(Y, Z)}{\partial Z^2} \right|_{Z=Z_0}$$

Has been recognized as far back as 1986 [Tierney, Kadane], but now we have the optimization tools to use this in nontrivial problems !

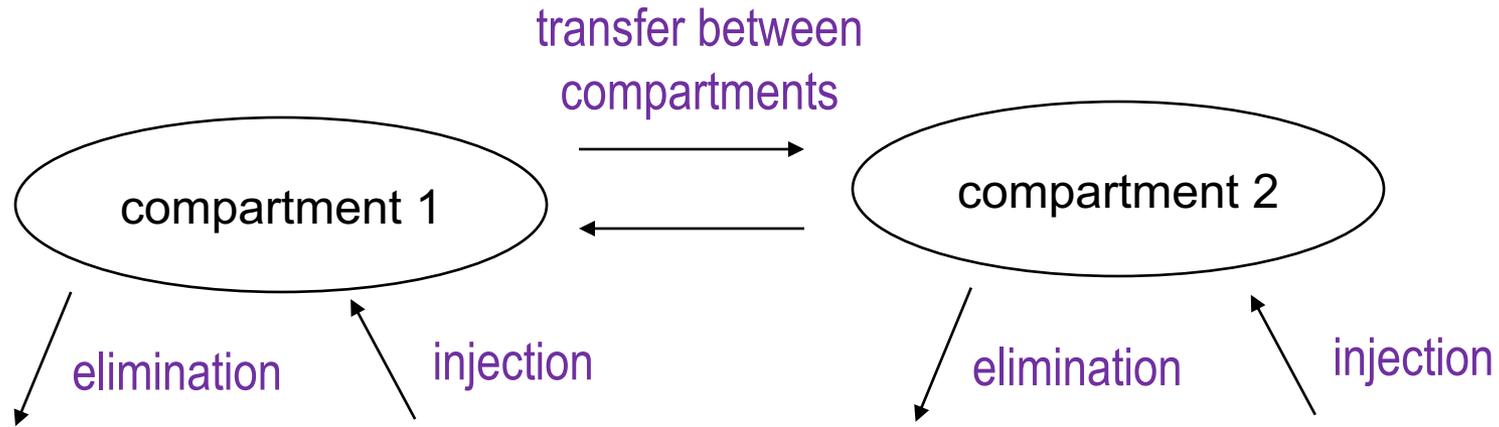
computed at maximum  $Z_0$

conditional means

$$E[Z|Y] = \int_{\mathbb{R}^n} Z e^{p(Z|Y)} dZ = \int_{\mathbb{R}^n} Z e^{p(Y, Z) - p(Y)} dZ \approx \dots = Z_0$$

$$\text{CoV}[Z|Y] = E[(Z - E[Z])(Z - E[Z])' | Y] \approx \dots = H(Y, Z_0)^{-1}$$

# Pharmacokinetic Model



$$C_i(t + 1) = (1 - K_{E_i})C_i(t) + \sum_j K_{ji}(C_j(t) - C_i(t)) + \frac{1}{V_i}u_i(t)$$

drug concentration in compartment  $i$  at time  $t + 1$  [moles/liter]

elimination rate

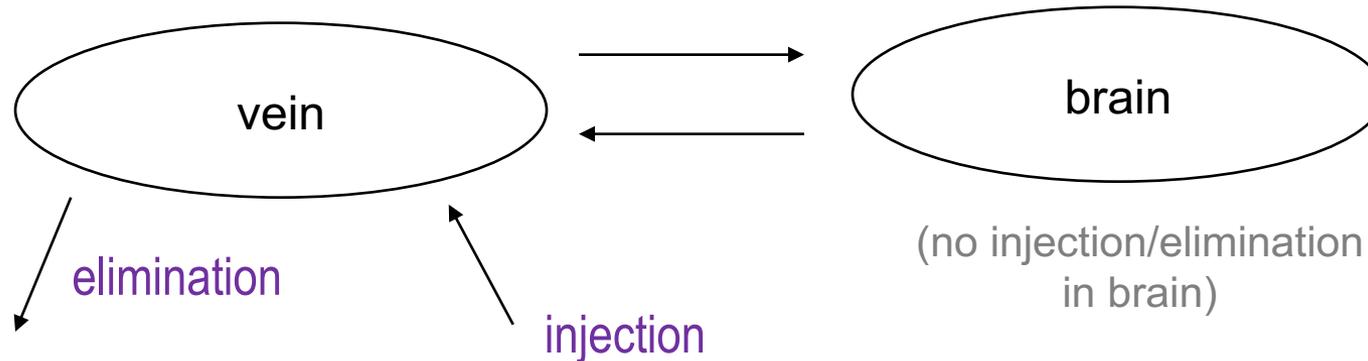
transfer rate between compartments  $i$  &  $j$  (Fick's law)

volume of compartment  $i$  [liter]

injection at time  $t$  [moles]

# Vancomycin Rat Model

Vancomycin is an antibiotic medication used to treat a number of bacterial infections



Vancomycin concentrations

$$C_{\text{vein}}(t + 1) = (1 - K_E)C_{\text{vein}}(t) + K_{BV}C_{\text{brain}}(t) + \frac{u_{\text{vein}}(t)}{V_{\text{vein}}}$$

$$C_{\text{brain}}(t + 1) = K_{IO}(C_{\text{vein}}(t) - C_{\text{brain}}(t))$$

typically small since  
 $\propto \frac{\text{brain volume}}{\text{vein volume}}$

(vein) injection  
at time  $t$   
[moles]

volume of vein  
compart. [liter]

meas. noise

$$\left. \begin{aligned} y_{\text{vein}}(t) &= C_{\text{vein}}(t) + w_{\text{vein}}(t) \\ y_{\text{brain}}(t) &= C_{\text{brain}}(t) + w_{\text{brain}}(t) \end{aligned} \right\}$$

measurements from E-AB (Electrochemical-aptamer based) sensors [Plaxco Lab]

- in-vivo
- time-resolution of second

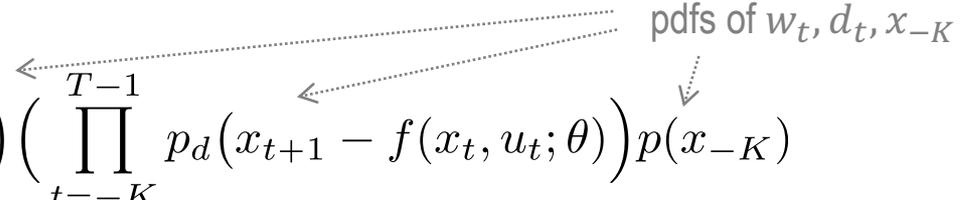
$$x_{t+1} = f(x_t, u_t; \theta) + d_t$$

$$y_t = g(x_t; \theta) + w_t$$

Observation: assuming  $x_{-K}, d_t, w_t$  independent, joint pdf of  $x_{-K}, d_t, y_t$  is easy to write

$$\left( \prod_{t=-K}^0 p_w(y_t - g(x_t; \theta)) \right) \left( \prod_{t=-K}^{T-1} p_d(x_{t+1} - f(x_t, u_t; \theta)) \right) p(x_{-K})$$

pdfs of  $w_t, d_t, x_{-K}$



Why? Proof by induction to compute

$$p(x_{t+1}, y_t \mid d_{-K}, \dots, d_t, y_{-K}, \dots, y_{t-1}, x_{-K}; \theta)$$

starting at  $t = -K$  & using Bayes rule for induction step

*Joint state-output pdf of a nonlinear system is easy to compute...*

# Vancomycin Rat Model

Vancomycin  
concentrations

$$C_{\text{vein}}(t+1) = (1 - K_E)C_{\text{vein}}(t) + \frac{u_{\text{vein}}(t)}{V_{\text{vein}}}$$
$$C_{\text{brain}}(t+1) = K_{IO}(C_{\text{vein}}(t) - C_{\text{brain}}(t))$$

(vein) injection  
at time  $t$   
[moles]

$$y_{\text{vein}}(t) = C_{\text{vein}}(t) + w_{\text{vein}}(t)$$
$$y_{\text{brain}}(t) = C_{\text{brain}}(t) + w_{\text{brain}}(t)$$

Defining

$$Y := \left\{ y_{\text{vein}}(1), y_{\text{brain}}(1), \dots, y_{\text{vein}}(T), y_{\text{brain}}(T) \right\} \quad \text{measurements}$$

$$Z := \left\{ K_E, 1/V_{\text{vein}}, K_{IO}, \quad \text{unknown parameters – large subject-2-subject variability} \right.$$
$$\left. C_{\text{brain}}(1), C_{\text{vein}}(1) \right\} \quad \text{initial conditions (known to be zero for our Vancomycin experiments)}$$

$$\theta := \left\{ \sigma_{\text{vein}}^2, \sigma_{\text{brain}}^2 \right\} \quad \text{unknown variances for measurement errors}$$

– large experiment-2-experiment variability

*Joint pdf  $p(Y, Z; \theta)$  is easy to compute...*

# Vancomycin Rat Model

## Defining

$$Y := \left\{ y_{\text{vein}}(1), y_{\text{brain}}(1), \dots, y_{\text{vein}}(T), y_{\text{brain}}(T) \right\} \quad \text{measurements}$$

$$Z := \left\{ K_E, 1/V_{\text{vein}}, K_{IO}, \quad \text{unknown parameters – large subject-2-subject variability} \right. \\ \left. C_{\text{brain}}(1), C_{\text{vein}}(1) \right\} \quad \text{initial conditions (known to be zero for our Vancomycin experiments)}$$

$$\theta := \left\{ \sigma_{\text{vein}}^2, \sigma_{\text{brain}}^2 \right\} \quad \text{unknown variances for measurement errors} \\ \text{– large experiment-2-experiment variability}$$

## From Laplace integration formulas ...

noise variance estimation:

$$\max_{\theta} p(Y; \theta) \propto \max_{\theta} \max_Z \frac{p(Y, Z; \theta)}{\det H(Y, Z; \theta)}$$

MLE to estimate  
for noise variances

optimization  
for Laplace integration

parameter estimation: arg-max from  
MLE estimate

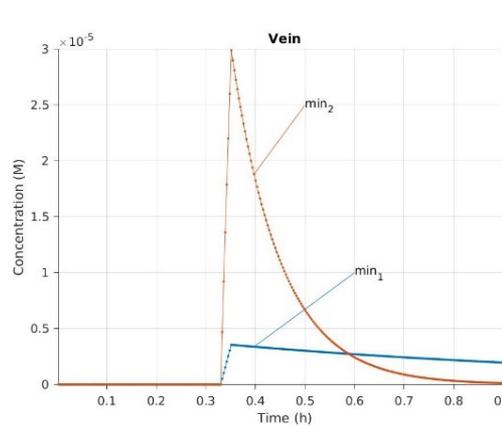
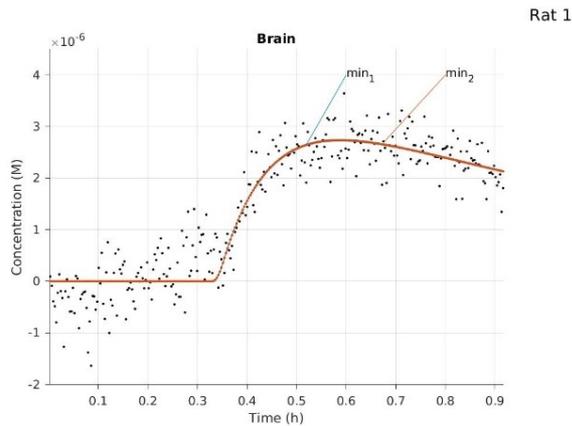
$$E[Z|Y] \approx Z_0 \\ \text{CoV}[Z|Y] \approx \frac{\partial^2 \log p(Y, Z)}{\partial Z^2} \Big|_{Z=Z_0}$$

a-posteriori parameter estimates  
and error covariances

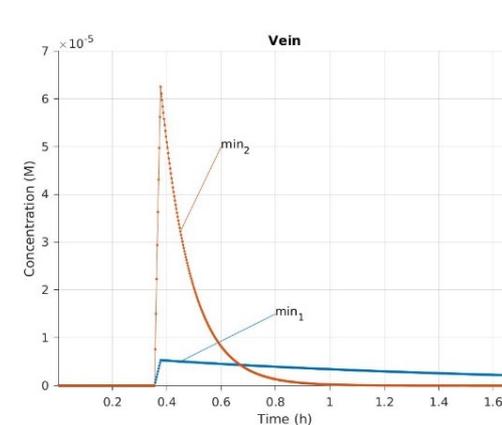
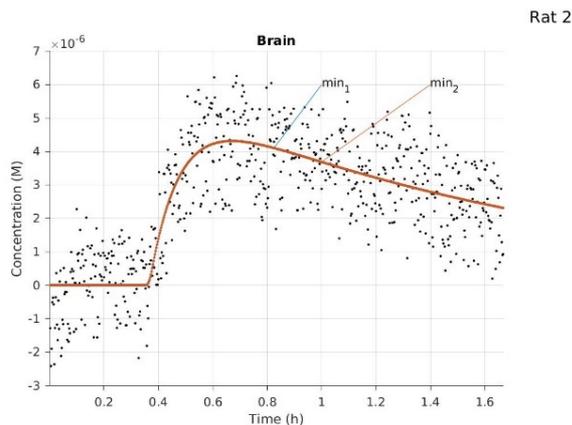
# Experimental results

Results based solely on measurements of brain concentration

individual 1



individual 2



concentration in brain  
(measurements  
& state estimates)

concentration in vein  
(state estimates  
for 2 global minima)

Just with brain  
measurements,  
optimization has 2  
isolated global  
maxima

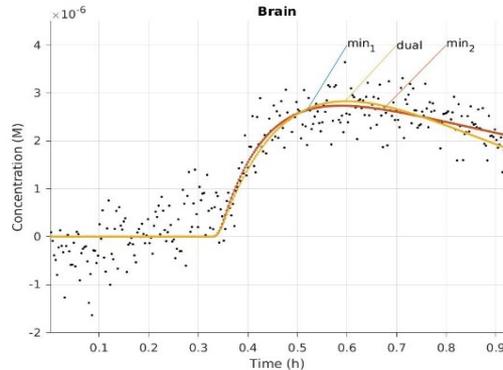


fundamental  
ambiguity in  
determining  
parameters &  
vein concentration

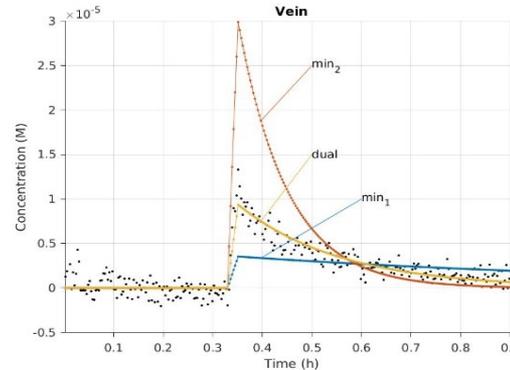
# Experimental results

Results based solely on measurements of brain concentration

individual 1



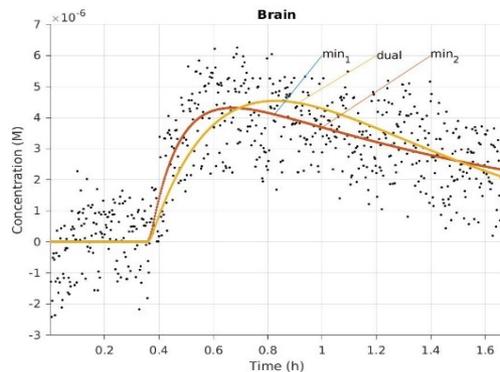
Rat 1



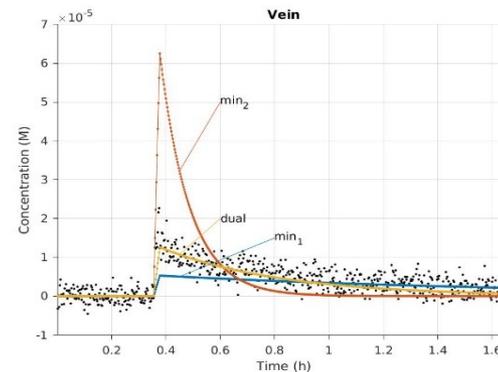
Single maximum  
with both brain &  
vein measurements

(previous maxima also  
shown in right-figures)

individual 2



Rat 2



concentration in brain  
(measurements  
& state estimates)

concentration in vein  
(measurement & state  
estimates)

Stochastic SIR model:

$$\begin{aligned} S(t+1) &= S(t) - \nu(t) && \text{susceptible} \\ I(t+1) &= I(t) + \nu(t) - \rho(t) && \text{infected} \\ R(t+1) &= R(t) + \rho(t) && \text{removed} \end{aligned}$$

$$\begin{aligned} \nu(t) &= \beta(t) \frac{I(t)}{N_0} S(t) + d_\nu(t) \\ \rho(t) &= \gamma(t) I(t) + d_\rho(t) \end{aligned}$$

daily new cases

daily new removals

time-varying infection/removal rates

stochastic disturbances

Measurement model:

$$\begin{aligned} y_\nu(t) &= \phi(t) \nu(t) + w_\nu(t) \\ y_D(t) &= \omega(t) I(t) + w_D(t) \end{aligned}$$

reported daily new cases

reported daily deaths

measurement noise

time-varying reporting rates

A-priori parameter drift model:

$$\beta(t), \gamma(t), \phi(t), \omega(t)$$

random walks with unknown variances

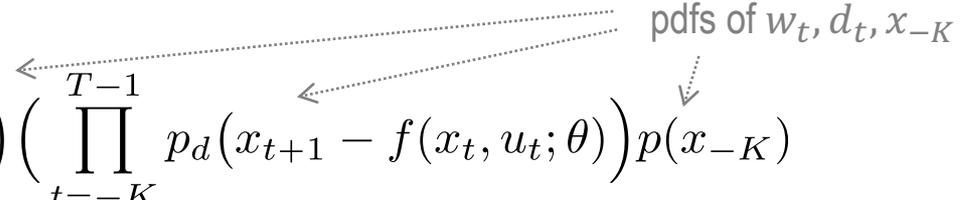
$$x_{t+1} = f(x_t, u_t; \theta) + d_t$$

$$y_t = g(x_t; \theta) + w_t$$

Observation: assuming  $x_{-K}, d_t, w_t$  independent, joint pdf of  $x_{-K}, d_t, y_t$  is easy to write

$$\left( \prod_{t=-K}^0 p_w(y_t - g(x_t; \theta)) \right) \left( \prod_{t=-K}^{T-1} p_d(x_{t+1} - f(x_t, u_t; \theta)) \right) p(x_{-K})$$

pdfs of  $w_t, d_t, x_{-K}$



Why? Proof by induction to compute

$$p(x_{t+1}, y_t \mid d_{-K}, \dots, d_t, y_{-K}, \dots, y_{t-1}, x_{-K}; \theta)$$

starting at  $t = -K$  & using Bayes rule for induction step

*Joint state-output pdf of a nonlinear system is easy to compute...*

Stochastic SIR model:

$$S(t+1) = S(t) - \nu(t) \quad \text{susceptible}$$

$$I(t+1) = I(t) + \nu(t) - \rho(t) \quad \text{infected}$$

$$R(t+1) = R(t) + \rho(t) \quad \text{removed}$$

$$\nu(t) = \beta(t) \frac{I(t)}{N_0} S(t) + d_\nu(t)$$

$$\rho(t) = \gamma(t) I(t) + d_\rho(t)$$

$$y_\nu(t) = \phi(t) \nu(t) + w_\nu(t)$$

$$y_D(t) = \omega(t) I(t) + w_D(t)$$

Defining

$$Y := \{y_\nu(1), y_D(1), \dots, y_\nu(T), y_D(T)\} \quad \text{past measurements}$$

$$Z := \{S(1), I(1), R(1), \beta(1), \gamma(1), \phi(1), \omega(1), \dots, \omega(T+P), \quad \text{past and future states}$$

$$y_\nu(T+1), y_D(T+1), \dots, y_\nu(T+P), y_D(T+P)\} \quad \text{future measurements}$$

*Joint pdf  $p(Y, Z)$  is easy to compute...*

## Defining

$$Y := \{y_\nu(1), y_D(1), \dots, y_\nu(T), y_D(T)\} \quad \text{past measurements}$$

$$Z := \left\{ S(1), I(1), R(1), \beta(1), \gamma(1), \phi(1), \omega(1), \dots, \omega(T+P), \right. \\ \left. y_\nu(T+1), y_D(T+1), \dots, y_\nu(T+P), y_D(T+P) \right\} \quad \begin{array}{l} \text{past and future states} \\ \text{future measurements} \end{array}$$

$$Z_0 := \arg \max_Z \log p(Y, Z)$$

## Identification:

$$\max_{\theta} p(Y; \theta) \propto \max_{\theta} \max_Z \frac{p(Y, Z; \theta)}{\det H(Y, Z; \theta)}$$

MLE to estimate  
unknown  
for noise/disturbance  
variances

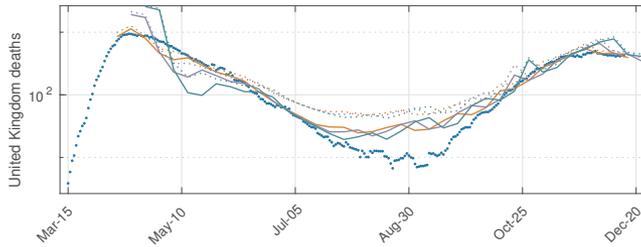
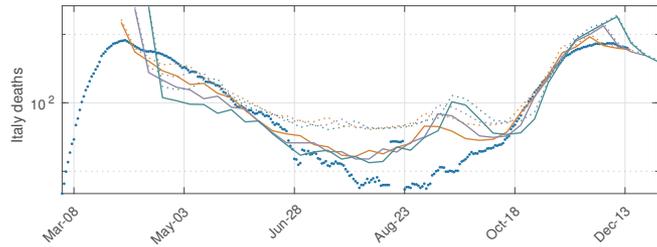
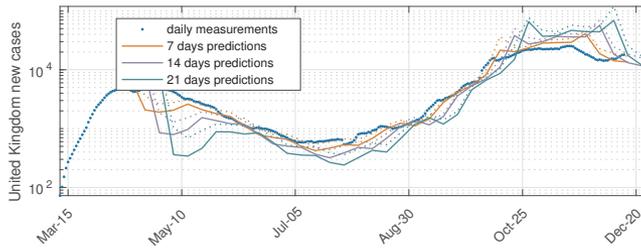
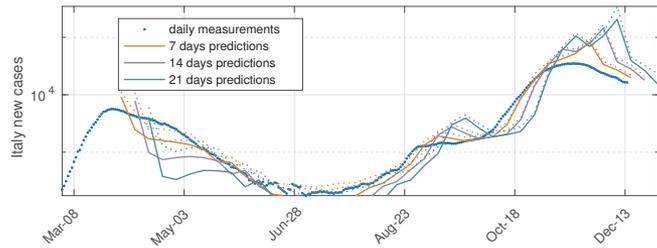
optimization  
for Laplace integration

## Forecasting:

$$E[Z|Y] \approx Z_0 \quad \leftarrow \text{arg-max from MLE estimate}$$
$$\text{CoV}[Z|Y] \approx \left. \frac{\partial^2 \log p(Y, Z)}{\partial Z^2} \right|_{Z=Z_0}$$

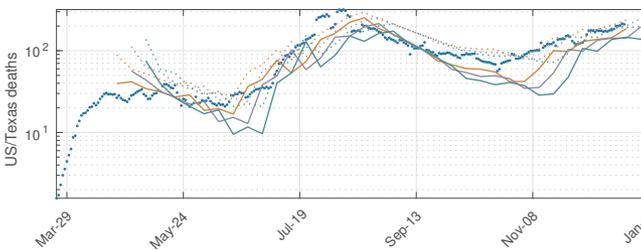
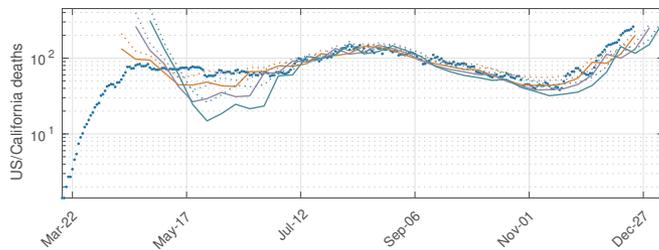
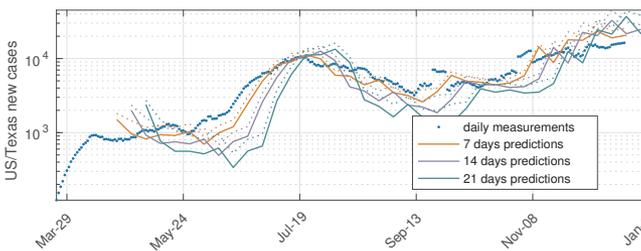
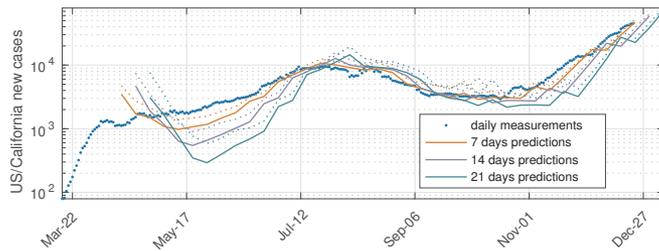
a-posteriori forecasts  
and error covariances

# Results



(a) Italy

(b) United Kingdom



(c) California, USA

(d) Texas, USA

Model not identifiable, but still possible to reliably compute

1. Maximum likelihood estimates for unknown random walk variances
2. 7-14-21 day Bayesian forecasts for measured variables

- Problem addressed:  $\min_u E_d [f(u, d)]$
- no closed form for expected value
  - willing to sacrifice accuracy for speed
  - error bounds

## Optimization-based bounds

1. Additive/multiplicative bounds for expected values
2. Laplace method for integration

## *What I did not talk about:*

- Numerical methods to solve min-max optimizations
  - primal-dual interior point methods for min-max equilibria
  - MATLAB toolbox (TensCalc)

## *Future/current work:*

- Which methods/bounds to use?
- Combination with MC methods

## *In collaboration with:*

*Raphael Chinchilla, Murat Erdal, Dr. Guosong Yang (UCSB)*

*Prof. Ramon Costa (Federal Univ. Rio Janeiro, Brazil)*

*Prof. Kevin Plaxco (UCSB)*

- Problem addressed:  $\min_u E_d [f(u, d)]$
- no closed form for expected value
  - willing to sacrifice accuracy for speed
  - error bounds

## Optimization-based bounds

1. Additive/multiplicative bounds for expected values
2. Laplace method for integration

## *What I did not talk about:*

- Numerical methods to solve min-max optimizations
  - primal-dual interior point methods for min-max equilibria
  - MATLAB toolbox (TensCalc)

## *Future/current work:*

- Which methods/bounds to use?
- Combination with MC methods

**THANKS!**

*In collaboration with:*

*Raphael Chinchilla, Murat Erdal, Dr. Guosong Yang (UCSB)*

*Prof. Ramon Costa (Federal Univ. Rio Janeiro, Brazil)*

*Prof. Kevin Plaxco (UCSB)*